Sparse Parameter Estimation: Compressed Sensing meets Matrix Pencil

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- Y. Chen and Y. Chi, Robust Spectral Compressed Sensing via Structured Matrix Completion, IEEE Trans. Information Theory, http://arxiv. org/abs/1304.8126
- Y. Li and Y. Chi, Off-the-Grid Line Spectrum Denoising and Estimation with Multiple Measurement Vectors, submitted, http://arxiv.org/abs/ 1408.2242

Sparse Fourier Analysis



In many sensing applications, one is interested in identification of a parametric signal:

$$x(\mathbf{t}) = \sum_{i=1}^{r} d_{i} e^{j2\pi \langle \mathbf{t}, \mathbf{f}_{i} \rangle}, \quad \mathbf{t} \in [\![n_{1}]\!] \times \ldots \times [\![n_{K}]\!]$$

 $(f_i \in [0,1)^K : \text{frequencies}, d_i : \text{amplitudes}, r : \text{model order})$

- Occam's razor: the number of modes r is small.
- Sensing with a minimal cost: how to identify the parametric signal model from a small subset of entries of x(t)?
- This problem has many (classical) applications in communications, remote sensing, and array signal processing.

Applications in Communications and Sensing

• Multipath channels: a (relatively) small number of strong paths.



• Radar Target Identification: a (relatively) small number of strong scatters.



Applications in Imaging

• Swap time and frequency:

$$z(t) = \sum_{i=1}^{r} d_i \delta(t - t_i)$$

- Applications in microscopy imaging and astrophysics.
- Resolution is limited by the point spread function of the imaging system





Something Old: Parametric Estimation

Exploring Physically-meaningful Constraints: *shift invariance* of the harmonic structure

$$x(\boldsymbol{t}-\boldsymbol{\tau}) = \sum_{i=1}^{r} d_{i} e^{j2\pi \langle \boldsymbol{t}-\boldsymbol{\tau}, \boldsymbol{f}_{i} \rangle} = \sum_{i=1}^{r} d_{i} e^{-j2\pi \langle \boldsymbol{\tau}, \boldsymbol{f}_{i} \rangle} e^{j2\pi \langle \boldsymbol{t}, \boldsymbol{f}_{i} \rangle}$$



- Prony's method: root-finding.
- SVD based approaches: ESPRIT [RoyKailath'1989], MUSIC [Schmidt'1986], matrix pencil [HuaSarkar'1990, Hua'1992].
- spectrum blind sampling [Bresler' 1996], finite rate of innovation [Vetterli' 2001], Xampling [Eldar' 2011].
- Pros: perfect recovery from (equi-spaced) O(r) samples
- Cons: sensitive to noise and outliers, usually require prior knowledge on the model order.

Something New: Compressed Sensing

Exploring Sparsity: Compressed Sensing [Candes and Tao'2006, Donoho'2006] capture the attributes (sparsity) of signals from a small number of samples.





• Discretize the frequency and assume a sparse representation over the discretized basis

$$f_i \in \mathcal{F} = \left\{\frac{0}{n_1}, \dots, \frac{n_1 - 1}{n_1}\right\} \times \left\{\frac{0}{n_2}, \dots, \frac{n_2 - 1}{n_2}\right\} \times \dots$$

- Pros: perfect recovery from $O(r \log n)$ random samples, robust against noise and outliers
- Cons: sensitive to gridding error

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Sensitivity to Basis Mismatch

- A toy example: $x(t) = e^{j2\pi f_0 t}$:
 - The success of CS relies on sparsity in the DFT basis.
 - Basis mismatch: Physics places f off grid by a frequency offset.
 - * Basis mismatch translates a sparse signal into an incompressible signal.



- Finer grid does not help, and one never estimates the true continuousvalued frequencies! [Chi, Scharf, Pezeshki, Calderbank 2011]

- Conventional approaches enforce physically-meaningful constraints, but not sparsity;
- Compressed sensing enforces sparsity, but not physically-meaningful constraints;
- Approach: We combine sparsity with physically-meaningful constraints, so that we can stably estimate the continuous-valued frequencies from a minimal number of time-domain samples.

revisit matrix pencil proposed for array signal processing

 revitalize matrix pencil by combining it with convex optimization



- Stack the signal $x(t) = \sum_{i=1}^{r} d_i e^{j2\pi \langle t, f_i \rangle}$ into a matrix $X \in \mathbb{C}^{n_1 \times n_2}$.
- The matrix X has the following Vandermonde decomposition:

$$X = Y \cdot \underbrace{D}_{\text{diagonal matrix}} \cdot Z^T.$$

Here,
$$D \coloneqq diag \{d_1, \dots, d_r\}$$
 and
 $Y \coloneqq \begin{bmatrix} 1 & 1 & \dots & 1 \\ y_1 & y_2 & \dots & y_r \\ \vdots & \vdots & \vdots & \vdots \\ y_1^{n_1-1} & y_2^{n_1-1} & \dots & y_r^{n_1-1} \end{bmatrix}, Z \coloneqq \begin{bmatrix} 1 & 1 & \dots & 1 \\ z_1 & z_2 & \dots & z_r \\ \vdots & \vdots & \vdots & \vdots \\ z_1^{n_2-1} & z_2^{n_2-1} & \dots & z_r^{n_2-1} \end{bmatrix}$
Vandemonde matrix Vandemonde matrix

where $y_i = \exp(j2\pi f_{1i})$, $z_i = \exp(j2\pi f_{2i})$, $f_i = (f_{1i}, f_{2i})$.

• Goal: We observe a *random subset of entries* of X, and wish to recover the missing entries.

Matrix Completion?

recall that
$$X = \underbrace{Y}_{Vandemonde} \cdot \underbrace{D}_{Vandemonde} \cdot \underbrace{Z}_{Vandemonde}^{T}$$
.

where $oldsymbol{D}\coloneqq ext{diag}\left\{ d_{1}, \cdots, d_{r}
ight\}$, and

- Quick observation: X can be a low-rank matrix with rank(X) = r.
- Question: can we apply *Matrix Completion* algorithms directly on X?

$$\begin{bmatrix} \sqrt{2} & ? & ? & \sqrt{2} & \sqrt{2} \\ ? & \sqrt{2} & ? & \sqrt{2} & \sqrt{2} \\ ? & ? & \sqrt{2} & \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} & \sqrt{2} & \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} & ? & \sqrt{2} & \sqrt{2} \end{bmatrix}$$

- Yes, but it yields sub-optimal performance.
 - It requires at least $r \max\{n_1, n_2\}$ samples.
- No, X is no longer low-rank if r > min (n₁, n₂)
 Note that r can be as large as n₁n₂

Given a data matrix X, Hua proposed the following matrix enhancement for two-dimensional frequency models [Hua 1992]:

• Choose two pencil parameters k_1 and k_2 ;



• An enhanced form X_e is an $k_1 \times (n_1 - k_1 + 1)$ block Hankel matrix :

$$\boldsymbol{X}_{e} = \begin{bmatrix} \boldsymbol{X}_{0} & \boldsymbol{X}_{1} & \cdots & \boldsymbol{X}_{n_{1}-k_{1}} \\ \boldsymbol{X}_{1} & \boldsymbol{X}_{2} & \cdots & \boldsymbol{X}_{n_{1}-k_{1}+1} \\ \vdots & \vdots & \vdots & \vdots \\ \boldsymbol{X}_{k_{1}-1} & \boldsymbol{X}_{k_{1}} & \cdots & \boldsymbol{X}_{n_{1}-1} \end{bmatrix},$$

where each block is a $k_2 \times (n_2 - k_2 + 1)$ Hankel matrix as follows

$$\boldsymbol{X}_{l} = \begin{bmatrix} x_{l,0} & x_{l,1} & \cdots & x_{l,n_{2}-k_{2}} \\ x_{l,1} & x_{l,2} & \cdots & x_{l,n_{2}-k_{2}+1} \\ \vdots & \vdots & \vdots & \vdots \\ x_{l,k_{2}-1} & x_{l,k_{2}} & \cdots & x_{l,n_{2}-1} \end{bmatrix}$$

- Choose pencil parameters $k_1 = \Theta(n_1)$ and $k_2 = \Theta(n_2)$, the dimensionality of X_e is proportional to $n_1n_2 \times n_1n_2$.
- The enhanced matrix can be decomposed as follows [Hua 1992]:

$$\boldsymbol{X}_{e} = \begin{bmatrix} \boldsymbol{Z}_{L} \\ \boldsymbol{Z}_{L} \boldsymbol{Y}_{d} \\ \vdots \\ \boldsymbol{Z}_{L} \boldsymbol{Y}_{d}^{k_{1}-1} \end{bmatrix} \boldsymbol{D} \begin{bmatrix} \boldsymbol{Z}_{R}, \boldsymbol{Y}_{d} \boldsymbol{Z}_{R}, \cdots, \boldsymbol{Y}_{d}^{n_{1}-k_{1}} \boldsymbol{Z}_{R} \end{bmatrix},$$

- Z_L and Z_R are Vandermonde matrices specified by z_1, \ldots, z_r , - $Y_d = diag[y_1, y_2, \cdots, y_r]$.
- The enhanced form $X_{
 m e}$ is low-rank.
 - $\operatorname{rank}(\boldsymbol{X}_{e}) \leq r$
 - − Spectral Sparsity ⇒ Low Rankness
- holds even with damping modes.



Enhanced Matrix Completion (EMaC)

• The natural algorithm is to find the enhanced matrix with the minimal rank satisfying the measurements:

 $\begin{array}{ll} \underset{\boldsymbol{M} \in \mathbb{C}^{n_1 \times n_2}}{\text{minimize}} & \operatorname{rank} \left(\boldsymbol{M}_{\mathrm{e}} \right) \\ \\ \text{subject to} & \boldsymbol{M}_{i,j} = \boldsymbol{X}_{i,j}, \forall (i,j) \in \Omega \end{array}$

where Ω denotes the sampling set.

• Motivated by Matrix Completion, we will solve its convex relaxation,

$$\begin{array}{ll} (\mathsf{EMaC}): & \underset{\boldsymbol{M} \in \mathbb{C}^{n_1 \times n_2}}{\operatorname{subject to}} & \|\boldsymbol{M}_{\mathsf{e}}\|_* \\ & \text{subject to} & \boldsymbol{M}_{i,j} = \boldsymbol{X}_{i,j}, \forall (i,j) \in \Omega \end{array}$$

where $\|\cdot\|_*$ denotes the nuclear norm.

• The algorithm is referred to as *Enhanced Matrix Completion (EMaC)*.

$$\begin{array}{ll} (\mathsf{EMaC}): & \underset{\boldsymbol{M} \in \mathbb{C}^{n_1 \times n_2}}{\operatorname{subject to}} & \|\boldsymbol{M}_{\mathsf{e}}\|_{\star} \\ & \text{subject to} & \boldsymbol{M}_{i,j} = \boldsymbol{X}_{i,j}, \forall (i,j) \in \Omega \end{array}$$

- existing MC result won't apply requires at least $\mathcal{O}(nr)$ samples
- **Question**: How many samples do we need?

$$\begin{bmatrix} ? & \sqrt{2} \\ \sqrt{2} & 7 & 7 & \sqrt{2} & \sqrt{2} & 7 & \sqrt{2} & \sqrt{2} & \sqrt{2} \\ \sqrt{2} & 7 & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} & \sqrt{2} & 7 & \sqrt{2} & \sqrt{2} & \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \\ \sqrt{2} & \sqrt{2} \end{bmatrix} \end{bmatrix}$$

• Define the 2-D Dirichlet kernel:

$$\mathcal{D}(k_1, k_2, f_1, f_2) \coloneqq \frac{1}{k_1 k_2} \left(\frac{1 - e^{-j2\pi k_1 f_1}}{1 - e^{-j2\pi f_1}} \right) \left(\frac{1 - e^{-j2\pi k_2 f_2}}{1 - e^{-j2\pi f_2}} \right),$$

• Define G_L and G_R as $r \times r$ Gram matrices such that

$$(\mathbf{G}_{\mathsf{L}})_{i,l} = \mathcal{D}(k_1, k_2, f_{1i} - f_{1l}, f_{2i} - f_{2l}),$$

$$(\mathbf{G}_{\mathsf{R}})_{i,l} = \mathcal{D}(n_1 - k_1 + 1, n_2 - k_2 + 1, f_{1i} - f_{1l}, f_{2i} - f_{2l}).$$



• Incoherence condition holds w.r.t. μ if

$$\sigma_{\min}(\boldsymbol{G}_L) \geq \frac{1}{\mu}, \quad \sigma_{\min}(\boldsymbol{G}_R) \geq \frac{1}{\mu}.$$

- Examples: $\mu = \Theta(1)$ under many scenarios:
 - Randomly generated frequencies;
 - (Mild) perturbation of grid points;
 - In 1D, let $k_1 \approx \frac{n_1}{2}$: well-separated frequencies (Liao and Fannjiang, 2014): $\Delta = \min_{i \neq j} |f_i - f_j| \gtrsim \frac{2}{n_1}$, which is about 2 times Rayleigh limits.



• Theorem [Chen and Chi, 2013] (Noiseless Samples) Let $n = n_1 n_2$. If all measurements are noiseless, then EMaC recovers X perfectly with high probability if

 $m > C \mu r \log^4 n.$

where C is some universal constant.

• Implications

- deterministic signal model, random observation;
- coherence condition μ only depends on the frequencies but not amplitudes.
- near-optimal within logarithmic factors: $\Theta(r \text{polylog}n)$.
- general theoretical guarantees for Hankel (Toeplitz) matrix completion.
 see applications in control, MRI, natural language processing, etc

Phase Transition



Figure 1: Phase transition diagrams where spike locations are randomly generated. The results are shown for the case where $n_1 = n_2 = 15$.

Robustness to Bounded Noise

Assume the samples are noisy $X = X^{o} + N$, where N is bounded noise:

$$\begin{array}{ll} (\mathsf{EMaC-Noisy}): & \underset{\boldsymbol{M} \in \mathbb{C}^{n_1 \times n_2}}{\operatorname{subject to}} & \left\| \boldsymbol{M}_{\mathsf{e}} \right\|_* \\ & \text{subject to} & \left\| \mathcal{P}_{\Omega} \left(\boldsymbol{M} - \boldsymbol{X} \right) \right\|_{\mathsf{F}} \leq \delta, \end{array}$$

• Theorem [Chen and Chi, 2013] (Noisy Samples) Suppose X° is a noisy copy of X that satisfies

$$\|\mathcal{P}_{\Omega}(\boldsymbol{X}-\boldsymbol{X}^{\mathrm{o}})\|_{F} \leq \delta.$$

Under the conditions of Theorem 1, the solution to EMaC-Noisy satisfies

$$\|\hat{\boldsymbol{X}}_{\mathrm{e}} - \boldsymbol{X}_{\mathrm{e}}\|_{F} \leq \left\{2\sqrt{n} + 8n + \frac{8\sqrt{2}n^{2}}{m}\right\}\delta$$

with probability exceeding $1 - n^{-2}$.

• Implications: The average entry inaccuracy is bounded above by $O(\frac{n}{m}\delta)$. In practice, EMaC-Noisy usually yields better estimate.

Singular Value Thresholding (Noisy Case)

• Several optimized solvers for Hankel matrix completion exist, for example [Fazel et. al. 2013, Liu and Vandenberghe 2009]

Algorithm 1 Singular Value Thresholding for EMaC

- 1: initialize Set $M_0 = X_e$ and t = 0.
- 2: repeat
- 3: 1) $\boldsymbol{Q}_t \leftarrow \mathcal{D}_{\tau_t}(\boldsymbol{M}_t)$ (singular-value thresholding)
- 4: 2) $M_t \leftarrow \text{Hankel}_{X_0}(Q_t)$ (projection onto a Hankel matrix consistent with observation)
- 5: 3) $t \leftarrow t + 1$
- 6: until convergence



Figure 2: dimension: 101×101 , r = 30, $\frac{m}{n_1 n_2} = 5.8\%$, SNR = 10dB.

• What if a constant portion of measurements are arbitrarily corrupted?



- Robust PCA approach [Candes et. al. 2011, Chandrasekaran et. al. 2011]
- Solve the following algorithm:

$$\begin{aligned} (\textbf{RobustEMaC}) : \min_{\boldsymbol{M}, \boldsymbol{S} \in \mathbb{C}^{n_1 \times n_2}} & \|\boldsymbol{M}_e\|_* + \lambda \|\boldsymbol{S}_e\|_1 \\ & \text{subject to} & (\boldsymbol{M} + \boldsymbol{S})_{i,l} = \boldsymbol{X}_{i,l}^{\text{corrupted}}, \quad \forall (i,l) \in \Omega \\ & \text{Page 22} \end{aligned}$$

• Theorem [Chen and Chi, 2013] (Sparse Outliers) Set $n = n_1 n_2$ and $\lambda = \frac{1}{\sqrt{m \log n}}$. Let the percent of corrupted entries $s \leq 20\%$ selected uniformly at random, then RobustEMaC recovers X with high probability if

$$m > C \mu r^2 \log^3 n,$$

where C is some universal constant.

• Implications:

- slightly more samples $m \sim \Theta(r^2 \log^3 n)$;
- robust to a constant portion of outliers: $s \sim \Theta(m)$;
- In summary, EMaC achieves robust recovery with respect to dense and sparse errors from a near-optimal number of samples.



Phase Transition for Line Spectrum Estimation

Fix the amount of corruption as 10% of the total number of samples:



Figure 3: Phase transition diagrams where spike locations are randomly generated. The results are shown for the case where n = 125.

- Cadzow's denoising with full observation: non-convex heuristic to denoise line spectrum data based on the Hankel form.
- Atomic norm minimization with random observation: recently proposed by [Tang et. al., 2013] for compressive line spectrum estimation off the grid.

$$\min_{\mathbf{s}} \| oldsymbol{s} \|_{\mathcal{A}} \quad ext{subject to} \quad \mathcal{P}_{\Omega}\left(oldsymbol{s}
ight) = \mathcal{P}_{\Omega}\left(oldsymbol{x}
ight),$$

where the atomic norm is defined as $\|\boldsymbol{x}\|_{\mathcal{A}} = \inf \left\{ \sum_{i} |d_{i}| | \boldsymbol{x}(t) = \sum_{i} d_{i} e^{j2\pi f_{i}t} \right\}.$

- Random signal model: if the frequencies are separated by 4 times the Rayleigh limit and the phases are random, then perfect recovery with $O(r \log r \log n)$ samples;
- no stability result with random observations;
- extendable to multi-dimensional frequencies [Chi and Chen, 2013], but the SDP characterization is more complicated [Xu et. al. 2013].

(Numerical) Comparison with Atomic Norm Minimization

Phase transition for 1D spectrum estimation: the phase transition for atomic norm minimization is very sensitive to the separation condition. The EMaC, in contrast, is insensitive to the separation.



Figure: phase transition for atomic norm minimization without separation (a), with separation (b); and EMaC without separation (c). The inclusion of separation doesn't change the phase transition of EMaC.

(Numerical) Comparison with Atomic Norm Minimization

Phase transition for 2D spectrum estimation: the phase transition for atomic norm minimization is very sensitive to the separation condition. The EMaC, in contrast, is insensitive to the separation. Here the problem dimension $n_1 = n_2 = 8$ is relatively small and the atomic norm minimization approach seems in favor despite of separation.



Figure: phase transition for atomic norm minimization without separation (a), with separation (b); and EMaC without separation (c). The inclusion of separation doesn't change the phase transition of EMaC.

Extension to Multiple Measurement Vectors Model

• When multiple snapshots available, it is possible to exploit the covariance structure to reduce the number of sensors. Without loss of generality, consider 1D:

$$x_{\ell}(t) = \sum_{i=1}^{r} d_{i,\ell} e^{j2\pi t f_i}, \quad t \in \{0, 1, \dots, n-1\}$$

where $x_{\ell} = [x_{\ell}(0), x_{\ell}(1), \dots, x_{\ell}(n-1)]^T$, $\ell = 1, \dots, L$.

• We assume the coefficients $d_{i,\ell} \sim \mathcal{CN}(0, \sigma_i^2)$, then the covariance matrix

$$\boldsymbol{\Sigma}$$
 = $\mathbb{E}\left[\boldsymbol{x}_{\ell}\boldsymbol{x}_{\ell}^{H}
ight]$ = toep (\boldsymbol{u})

is a PSD block Toeplitz matrix with rank(Σ) = r.

• The frequencies can be estimated without separation from u using ℓ_1 minimization with nonnegative constraints [Donoho and Tanner' 2005].

• Observation pattern: Instead of random observations, we assume deterministic observation pattern Ω over a (minimum) sparse ruler for all snapshots:

$$\boldsymbol{y}_{\ell} = \mathbf{x}_{\Omega,\ell} = \{x_{\ell}(t), \quad t \in \Omega\}, \quad \ell = 1, \dots, L.$$

- Sparse ruler in 1D: for $\Omega \in \{0, \dots, n-1\}$
 - Define the difference set:

$$\Delta = \{ |i - j|, \quad \forall i, j \in \Omega \}$$

- Ω is called a length-*n* sparse ruler if $\Delta = \{0, \ldots, n-1\}$.
- Examples:
 - * when n = 21, $\Omega = \{0, 1, 2, 6, 7, 8, 17, 20\}$.
 - * nested arrays, co-prime arrays [Pal and Vaidyanathan' 2010, 2011]
 - * minimum sparse rulers [Redei and Renyi, 1949]
- roughly $|\Omega| = O(\sqrt{n})$.

• Consider the observation on $\Omega = \{0, 1, 2, 5, 8\}$,

$$\mathbb{E} \left[\boldsymbol{y}_{\ell} \boldsymbol{y}_{\ell}^{H} \right] = \mathbb{E} \begin{bmatrix} x_{\ell}(0) \\ x_{\ell}(1) \\ x_{\ell}(2) \\ x_{\ell}(5) \\ x_{\ell}(8) \end{bmatrix} \begin{bmatrix} x_{\ell}^{H}(0) \ x_{\ell}^{H}(1) \ x_{\ell}^{H}(2) \ x_{\ell}^{H}(5) \ x_{\ell}^{H}(8) \end{bmatrix}$$
$$= \begin{bmatrix} u_{0} \ u_{1}^{H} \ u_{2}^{H} \ u_{1}^{H} \ u_{2}^{H} \ u_{1}^{H} \ u_{4}^{H} \ u_{7}^{H} \\ u_{2} \ u_{1} \ u_{0} \ u_{3}^{H} \ u_{6}^{H} \\ u_{5} \ u_{4} \ u_{3} \ u_{0} \ u_{3}^{H} \\ u_{8} \ u_{7} \ u_{6} \ u_{3} \ u_{0} \end{bmatrix}$$

• which gives the exact full covariance matrix $\Sigma = toep(u)$ in the absence of noise and an infinite number of snapshots.

• In practice, measurements will be noisy with a finite number of snapshots:

$$\boldsymbol{y}_{\ell} = \boldsymbol{x}_{\Omega,\ell} + \boldsymbol{w}_{\ell}, \quad \ell = 1, \dots, L,$$

where $\boldsymbol{w}_\ell \sim \mathcal{CN}(\sigma^2, \boldsymbol{I})$.

- **Two-step covariance estimation**:
 - formulate the sample covariance matrix of y_ℓ :

$$\boldsymbol{\Sigma}_{\Omega,L} = \frac{1}{L} \sum_{\ell=1}^{L} \boldsymbol{y}_{\ell} \boldsymbol{y}_{\ell}^{H};$$

- determine the Teoplitz covariance matrix with SDP:

$$\hat{\boldsymbol{u}} = \underset{\boldsymbol{u}: \text{toep}(\boldsymbol{u}) \geq 0}{\operatorname{argmin}} \ \frac{1}{2} \left\| \mathcal{P}_{\Omega}(\text{toep}(\boldsymbol{u})) - \boldsymbol{\Sigma}_{\Omega,L} \right\|_{F}^{2} + \lambda \mathsf{Tr}(\text{toep}(\boldsymbol{u})),$$

where λ is some regularization parameter.

• Theorem [Li and Chi] Let u^* be the ground truth. Set

$$\lambda \ge C \max\left\{\sqrt{\frac{r\log(Ln)}{L}}, \frac{r\log(Ln)}{L}\right\} \|\boldsymbol{\Sigma}_{\Omega}^{\star}\|$$

with $\Sigma_{\Omega}^{\star} = \mathbb{E}[\boldsymbol{y}_{m}\boldsymbol{y}_{m}^{H}]$ for some constant C, then with probability at least $1 - L^{-1}$, the solution satisfies

$$\frac{1}{\sqrt{n}} \|\hat{\boldsymbol{u}} - \boldsymbol{u}^{\star}\|_{F} \le 16\lambda\sqrt{r}$$

if Ω is a complete sparse ruler.

- Remark:
 - $-\frac{1}{\sqrt{n}}\|\hat{\boldsymbol{u}}-\boldsymbol{u}^{\star}\|_{F}$ is small as soon as $L \gtrsim O(r^{2}\log n)$;
 - As the rank r (as large as n) can be larger than $|\Omega|$ (as small as \sqrt{n}), this allows frequency estimation even the snapshots cannot be recovered.

The algorithm also applies to other observation patterns, e.g. random. Setting: n = 64, L = 400, $|\Omega| = 5$, and r = 6.



- Sparse parameter estimation is possible leveraging shift-invariance structures embedded in matrix pencil with recent matrix completion techniques;
- Fundamental performance is determined by the proximity of the frequencies measured by the conditioning number of the Gram matrix formed by the sampling the Dirichlet kernel;
- Recovering more lines than the number of sensors is made possible by exploiting the second-order statistics;
- **Future work:** how to compare conventional algorithms with those of convex optimization?

Publications available on arXiv:

- Robust Spectral Compressed Sensing via Structured Matrix Completion, IEEE Trans. Information Theory, http://arxiv.org/abs/1304.8126
- Off-the-Grid Line Spectrum Denoising and Estimation with Multiple Measurement Vectors, submitted, http://arxiv.org/abs/1408.2242

Thank You! Questions?